## Practice Questions for Midterm 10/36-702

(1) Let  $X_1, \ldots, X_n \sim \text{Unif}(0, 1)$ . Compute the bias and variance of the histogram density estimator with binwidth h for this distribution. Show that the optimal value of h is h = 1.

(2) Let  $X_1, \ldots, X_n \sim P$  where p has a density p on  $\mathbb{R}$ . Assume that p(x) > 0 for each  $x \in \mathbb{R}$ . Given  $c_1, \ldots, c_k \in \mathbb{R}$ , the population k-means risk is

$$R(k) = \inf_{c_1,\dots,c_k} \mathbb{E}\left(\min_{j=1,\dots,k} |X - c_j|^2\right).$$

Show that R(k) is strictly decreasing in k.

(3) Let  $(X_1, Y_1), \ldots, (X_n, Y_n)$  be iid. Suppose that  $X_1, \ldots, X_n \sim P$  has a density p on [0, 1] where  $0 < c \le p(x) \le C < \infty$  for all  $x \in [0, 1]$ . Assume that the density p is known. Assume that

$$Y_i = m(X_i) + \epsilon_i$$

where  $\epsilon_1, \ldots, \epsilon_n$  are iid with mean 0 and variance  $\sigma^2$ . Assume that m, m', m'', m''', p, p', p'', p''' are bounded and continuous functions. Let  $x \in (0, 1)$  and define

$$\widehat{m}(x) = \frac{\frac{1}{n} \sum_{i=1}^{n} Y_i \frac{1}{h} K\left(\frac{x - X_i}{h}\right)}{p(x)}$$

where K is a smooth, symmetric, kernel with bounded support. Show that

$$\mathbb{E}[\widehat{m}(x)] = m(x) + Ch^2 + O(h^3).$$

(4) Let  $(X_1, Y_1), \ldots, (X_n, Y_n)$  be iid. Suppose that  $Y_i \in \{0, 1\}$  and  $X_i \in [0, 1]$ . Let  $\theta = P(Y_i = 1)$ . Assume that  $0 < \theta < 1$ . Suppose that

$$X_i \mid Y_i = 1 ~\sim~ p_1$$

and

$$X_i \mid Y_i = 0 ~\sim ~ p_0$$

where  $p_0$  and  $p_1$  are densities on [0, 1]. Assume that, for some constants, c and C,

$$0 < c \le p_j(x) \le C < \infty$$

for all  $x \in [0, 1]$  and j = 0, 1.

Let  $\hat{p}_0$  be an estimate of  $p_0$  and let  $\hat{p}_1$  be an estimate of  $p_1$ . Define

$$\widehat{h}(x) = \begin{cases} 1 & \text{if } \widehat{m}(x) \ge 1/2\\ 0 & \text{if } \widehat{m}(x) < 1/2 \end{cases}$$

where

$$\widehat{m}(x) = \frac{\widehat{\theta}\,\widehat{p}_1(x)}{\widehat{\theta}\,\widehat{p}_1(x) + (1 - \widehat{\theta})\,\widehat{p}_0(x)}$$

 $\widehat{\theta} = n^{-1} \sum_{i=1}^{n} Y_i,$ 

Suppose that

$$\sup_{x} |\widehat{p}_{0}(x) - p_{0}(x)| \xrightarrow{P} 0, \quad \text{and} \quad \sup_{x} |\widehat{p}_{1}(x) - p_{1}(x)| \xrightarrow{P} 0$$

Show that

$$\mathbb{P}(Y \neq \widehat{h}(X)) - \mathbb{P}(Y \neq h_*(X)) \xrightarrow{P} 0$$

as  $n \to \infty$ , where  $h_*$  is the Bayes classifier, and  $\mathbb{P}$  is probability with respect to X and Y, but not with respect to  $\hat{h}$ .

(5) Let p be a bounded continuous density defined on a bounded subset  $S \subset \mathbb{R}$ . Assume further that p has bounded, continuous first and second derivatives. Let  $Y_1, \ldots, Y_n \sim p$  and let

$$\widehat{p}_h(y) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h} K\left(\frac{y-Y_i}{h}\right).$$

Let  $p_h(x) = \mathbb{E}[\widehat{p}_h(x)].$ 

(a) Show that, for any t > 0,  $\mathbb{P}(|\hat{p}_h(x) - p_h(x)| > t) \to 0$  as long as  $nh \to \infty$ .

(b) Let  $C_h = \{x : p_h(x) > \lambda\}$  and let  $\widehat{C}_h = \{x : \widehat{p}_h(x) > \lambda\}$ . Show that  $\widehat{C}_h$  is a consistent estimator of  $C_h$  in the following sense: (i) if  $p_h(x) > \lambda$  then  $\mathbb{P}(x \in \widehat{C}_h) \to 1$  and (ii) if  $p_h(x) < \lambda$  then  $\mathbb{P}(x \notin \widehat{C}_h) \to 1$ .

(6) Let  $Y_i = \beta^T X_i + \epsilon_i$  where  $Y_i \in \mathbb{R}$ ,  $X_i \in \mathbb{R}^d$  and  $\epsilon_i \sim N(0, \sigma^2)$ . Recall that the ridge estimator is

$$\widehat{\beta} = (\mathbb{X}^T \mathbb{X} + \lambda I)^{-1} \mathbb{X}^T Y,$$

where  $\mathbb{X} \in \mathbb{R}^{n \times d}$ ,  $Y = (Y_1, \ldots, Y_n)$  and  $\lambda \geq 0$ . Find  $\mathbb{E}[\widehat{\beta}|X_1, \ldots, X_n]$  and  $\operatorname{Var}[\widehat{\beta}|X_1, \ldots, X_n]$ . Show that  $\operatorname{Var}[\widehat{\beta}|X_1, \ldots, X_n] \to 0$  as  $\lambda \to \infty$ . Show that the bias tends to 0 as  $\lambda \to 0$  if d < n.

(7) Let  $(X, Y) \sim P$ , and consider predicting the value of Y from X. That is, consider choosing a function f to minimize

$$\mathbb{E}\big[\big(Y - f(X)\big)^2\big].$$

Show that the function minimizing this is given by

$$f(x) = \frac{\int y \cdot p_{X,Y}(x,y) \, dy}{p_X(x)},$$

where  $p_{X,Y}$  is the joint density of (X, Y), and  $p_X$  is the density of X.