# Sequential Nonparametric Two-sample Testing by Betting

# IISA Conference, 2022



Shubhanshu Shekhar



Aaditya Ramdas

Department of Statistics and Data Science Carnegie Mellon University

# Sequential Two-Sample Testing

+ Given a stream of paired observations on  $\mathcal{X}\times\mathcal{X}$ 

 $(X_1, Y_1), (X_2, Y_2), \ldots \sim P_X \times P_Y$  i.i.d.,

• decide between the hypotheses:

$$H_0: P_X = P_Y$$
 and  $H_1: P_X \neq P_Y$ .



For  $\alpha \in (0, 1)$ , construct a level- $\alpha$  sequential test of power one.

- Under  $H_0$ : continue forever w.p.  $\geq 1 \alpha$ .
- Under *H*<sub>1</sub>: stop sampling, and reject the null as soon as possible.

#### Batch Two-Sample Testing

- Here, we have batches of observations:  $(X_1, \ldots, X_n)$  and  $(Y_1, \ldots, Y_m)$  drawn i.i.d.from  $P_X$  and  $P_Y$  respectively.
- A popular class of batch tests based on statistical distances  $d: \mathcal{P}(\mathcal{X}) \times \mathcal{P}(\mathcal{X}) \to \mathbb{R}.$ 
  - Define a test statistic  $T_{n,m} = d(\widehat{P}_{X,n}, \widehat{P}_{Y,m}).$
  - Reject the null, if  $T_{n,m}$  is large.
- + E.g.,  $\chi^2$ -test, Kolmogorov-Smirnov (KS) test, kernel-MMD test.
- · Theoretical and empirical properties have been well studied.
- No such general framework for constructing sequential two-sample tests of power one.

#### Prior Sequential Nonparametric Tests of Power One

- Darling & Robbins (1968): based on time-uniform DKW inequalitites for univariate observations.
- Balsubramani & Ramdas (2016): based on a confidence sequence (CS) for linear-time kernel-MMD statistic
- Lheritier & Cazals (2017): based on sequential binary classifiers
- Howard & Ramdas (2021): based on CSs using forward supermartingales
- Manole & Ramdas (2021): based on CSs using reverse submartingales.

All existing methods either have strong theoretical guarantees or good empirical performance; but not both.

All existing methods either have strong theoretical guarantees or good empirical performance; but not both.

#### This Talk

- A fundamentally new framework for designing powerful sequential two-sample tests.
- We take the perspective of a fictitious bettor, repeatedly betting on the observations to disprove the null.
  - **Constraints:** The bets must be fair under *H*<sub>0</sub>, and the bettor cannot borrow money.
- The gain in the bettor's wealth (i.e.,  $W_t/W_0$ ) is a measure of evidence collected against the null.

#### The Betting Game

Bettor begins with an initial wealth,  $W_0 = 1$ .

For t = 1, 2, ...:

- Bettor selects a function  $g_t : \mathcal{X} \to [-1/2, 1/2]$ .
  - defines a fair payoff function under  $H_0$ ,  $h_t(x,y) = g_t(x) - g_t(y)$ .
- Bettor chooses a fraction,  $\lambda_t \in [0, 1]$ , of his current wealth,  $W_{t-1}$ , to gamble.
- The next paired observation,  $(X_t, Y_t)$ , is revealed.
- Bettor's wealth is updated as follows:

$$W_t = W_{t-1} \times (1 - \lambda_t) + W_{t-1}\lambda_t (1 + h_t(X_t, Y_t))$$
$$= W_0 \times \prod_{i=1}^t \left( 1 + \lambda_i (g_i(X_i) - g_i(Y_i)) \right)$$

### From Betting to Sequential Testing

Under  $H_0$ , we have  $\mathbb{E}[g_t(X_t) - g_t(Y_t)|\mathcal{F}_{t-1}] = 0$ . Hence,  $\{W_t : t \ge 0\}$  is a test martingale — a non-negative martingale with an initial value 1.

# From Betting to Sequential Testing

Under  $H_0$ , we have  $\mathbb{E}[g_t(X_t) - g_t(Y_t)|\mathcal{F}_{t-1}] = 0$ . Hence,  $\{W_t : t \ge 0\}$  is a test martingale — a non-negative martingale with an initial value 1.

#### Ville's Inequality (1939)

For any test martingale  $\{W_t : t \ge 0\}$  and an  $\alpha \in (0, 1]$ , we have

$$\mathbb{P}\left(\exists t \geq 0 : W_t \geq \frac{1}{\alpha}\right) \leq \alpha.$$

# JEAN VILLE Étude critique de la notion de collectif

Thèses de l'entre-deux-guerres, 1939



## From Betting to Sequential Testing

Under  $H_0$ , we have  $\mathbb{E}[g_t(X_t) - g_t(Y_t)|\mathcal{F}_{t-1}] = 0$ . Hence,  $\{W_t : t \ge 0\}$  is a test martingale — a non-negative martingale with an initial value 1.

#### Ville's Inequality (1939)

For any test martingale  $\{W_t : t \ge 0\}$  and an  $\alpha \in (0, 1]$ , we have

$$\mathbb{P}\left(\exists t \geq 0 : W_t \geq \frac{1}{\alpha}\right) \leq \alpha.$$

• Define the test (i.e., a stopping time):

$$\tau \coloneqq \min\{t \ge 1 : W_t \ge 1/\alpha\}.$$

• For arbitrary (predictable) sequences  $\{(g_t, \lambda_t) : t \ge 1\}$ , Ville's inequality implies

$$\mathbb{P}(\tau < \infty) \le \alpha$$
, under  $H_0$ .

Faster growth of  $W_t \Rightarrow$  Stronger statistical properties of  $\tau$ 

Faster growth of  $W_t \Rightarrow$  Stronger statistical properties of  $\tau$ 

· Consistency.

 $\mathbb{P}\left(\exists n \geq 1 : W_n \geq 1/\alpha\right) = 1 \quad \Rightarrow \quad \mathbb{P}(\tau < \infty) = 1.$ 

Faster growth of  $W_t \Rightarrow$  Stronger statistical properties of  $\tau$ 

 $\cdot$  Consistency.

$$\mathbb{P}\left(\exists n \geq 1 : W_n \geq 1/\alpha\right) = 1 \quad \Rightarrow \quad \mathbb{P}(\tau < \infty) = 1.$$

· Exponential consistency.

$$\liminf_{n\to\infty}\frac{-1}{n}\log\left(\mathbb{P}(W_n<1/\alpha)\right)>0\quad\Rightarrow\quad\liminf_{n\to\infty}\frac{-1}{n}\log\left(\mathbb{P}(\tau>n)\right)>0.$$

Faster growth of  $W_t \Rightarrow$  Stronger statistical properties of  $\tau$ 

• Consistency.

$$\mathbb{P}\left(\exists n \geq 1 : W_n \geq 1/\alpha\right) = 1 \quad \Rightarrow \quad \mathbb{P}(\tau < \infty) = 1.$$

· Exponential consistency.

$$\liminf_{n\to\infty}\frac{-1}{n}\log\left(\mathbb{P}(W_n<1/\alpha)\right)>0\quad\Rightarrow\quad\liminf_{n\to\infty}\frac{-1}{n}\log\left(\mathbb{P}(\tau>n)\right)>0.$$

· Finite Expected Stopping Time.

$$\sum_{n\geq 0} \mathbb{P}\left(W_n < \frac{1}{\alpha}\right) < \infty \quad \Rightarrow \quad \mathbb{E}[\tau] = \sum_{n=0}^{\infty} \mathbb{P}\left(\tau > n\right) < \infty.$$

#### Summary so far

- We defined a sequential test:  $\tau = \min\{t \ge 1 : W_t \ge 1/\alpha\}.$
- { $W_t : t \ge 1$ } is the wealth of a fictitious bettor, betting on the observations in a repeated game with  $W_0 = 1$ .
- Under  $H_0$ , for arbitrary predictable  $\{(g_t, \lambda_t) : t \ge 1\}$ , we have  $\mathbb{P}(\tau < \infty) \le \alpha$ .
- **Under**  $H_1$ , statistical properties of  $\tau$  depend on how quickly  $W_t$  grows to infinity.
  - this depends strongly on the choice of  $\{(\lambda_t, g_t) : t \ge 1\}$ .
- Rest of the talk: A principled approach for selecting  $\{(\lambda_t, g_t) : t \ge 1\}.$

- $\cdot$  Step 1: Select an appropriate function class  ${\cal G}$ 
  - Or equivalently, an Integral Probability Metric (IPM)
- Step 2: Design an "Oracle Test"
  - Uses terms,  $g^*$  and  $\lambda^*$ , depending on the unknown  $P_X$  and  $P_Y$
- **Step 3:** Design a practical sequential test
  - Uses a sequence of predictable estimates of  $g^*$  and  $\lambda^*$

- For simplicity, we assume that  $\mathcal G$  consists of functions taking values in [-1/2, 1/2].
- Can define

$$d_{\mathcal{G}}(P_{X}, P_{Y}) = \max_{g \in \mathcal{G}} \mathbb{E}_{P_{X}}[g(X)] - \mathbb{E}_{P_{Y}}[g(Y)],$$
(1)

which is a metric if  $\mathcal{G}$  is rich enough.

- For simplicity, we assume that  $\mathcal{G}$  consists of functions taking values in [-1/2, 1/2].
- Can define

$$d_{\mathcal{G}}(P_X, P_Y) = \max_{g \in \mathcal{G}} \mathbb{E}_{P_X}[g(X)] - \mathbb{E}_{P_Y}[g(Y)], \tag{1}$$

which is a metric if  $\mathcal{G}$  is rich enough.

• Witness function

$$g^* \in \underset{g \in \mathcal{G}}{\arg \max} \ \mathbb{E}_{P_{X}}[g(X)] - \mathbb{E}_{P_{Y}}[g(Y)].$$
(2)

- For simplicity, we assume that  $\mathcal{G}$  consists of functions taking values in [-1/2, 1/2].
- Can define

$$d_{\mathcal{G}}(P_X, P_Y) = \max_{g \in \mathcal{G}} \mathbb{E}_{P_X}[g(X)] - \mathbb{E}_{P_Y}[g(Y)], \tag{1}$$

which is a metric if  $\mathcal{G}$  is rich enough.

Witness function

$$g^* \in \underset{g \in \mathcal{G}}{\arg \max} \mathbb{E}_{P_{X}}[g(X)] - \mathbb{E}_{P_{Y}}[g(Y)].$$
(2)

•  $g^*$  provides the maximum contrast between  $P_X$  and  $P_Y$ 

- For simplicity, we assume that  $\mathcal{G}$  consists of functions taking values in [-1/2, 1/2].
- Can define

$$d_{\mathcal{G}}(P_X, P_Y) = \max_{g \in \mathcal{G}} \mathbb{E}_{P_X}[g(X)] - \mathbb{E}_{P_Y}[g(Y)], \tag{1}$$

which is a metric if  $\mathcal{G}$  is rich enough.

Witness function

$$g^* \in \underset{g \in \mathcal{G}}{\arg \max} \mathbb{E}_{P_{X}}[g(X)] - \mathbb{E}_{P_{Y}}[g(Y)].$$
(2)

- $\cdot g^*$  provides the maximum contrast between  $P_X$  and  $P_Y$
- If  $P_X = P_Y$ , then  $g^*$  is an arbitrary element of  $\mathcal{G}$

$$W_t^* = W_{t-1}^* \times (1 + \lambda^* (g^*(X_t) - g^*(Y_t))),$$

$$W_t^* = W_{t-1}^* \times (1 + \lambda^* (g^*(X_t) - g^*(Y_t))),$$

• where  $\lambda^*$  is the **log-optimal betting fraction**:

$$\lambda^* \in \underset{\lambda \in (-1,1)}{\operatorname{arg max}} \mathbb{E}\left[\log(1 + \lambda(g^*(X) - g^*(Y)))\right].$$

A New Interpretation of Information Rate  $_{reproduced \ with \ permission \ of \ AT\&T}$ 

By J. L. Kelly, jr.

(Manuscript received March 21, 1956)



$$W_t^* = W_{t-1}^* \times (1 + \lambda^* (g^*(X_t) - g^*(Y_t))),$$

• where  $\lambda^*$  is the **log-optimal betting fraction**:

$$\lambda^* \in \underset{\lambda \in (-1,1)}{\operatorname{arg\,max}} \mathbb{E}\left[\log(1 + \lambda(g^*(X) - g^*(Y)))\right].$$

• Define the 'oracle test':  $\tau^* = \min \{t \ge 1 : W_t^* \ge \frac{1}{\alpha}\}.$ 

$$W_t^* = W_{t-1}^* \times (1 + \lambda^* (g^*(X_t) - g^*(Y_t))),$$

• where  $\lambda^*$  is the **log-optimal betting fraction**:

$$\lambda^* \in \underset{\lambda \in (-1,1)}{\arg \max} \mathbb{E}\left[\log(1 + \lambda(g^*(X) - g^*(Y)))\right].$$

- Define the 'oracle test':  $\tau^* = \min \{t \ge 1 : W_t^* \ge \frac{1}{\alpha}\}.$
- The test  $\tau^*$  is exponentially consistent, and has a finite expected stopping time.

#### Step 3 – Practical Test

- $\cdot g^*$  and  $\lambda^*$  in  $\tau^*$  are not known  $\Rightarrow$  Use data-driven estimates.
- A prediction strategy  $(A_P)$  to select  $\{g_t : t \ge 1\} \approx g^*$ .
  - Specific choice of  $\mathcal{A}_{P}$  will depend on  $\mathcal{G}$ .
- A betting strategy  $(\mathcal{A}_B)$  to select  $\{\lambda_t : t \ge 1\} \approx \lambda^*$ .
  - Existing methods, such as Online Newton Step (ONS), are sufficient for our purposes.
- Construct the wealth process

$$W_t = W_{t-1} \times \left(1 + \lambda_t (g_t(X_t) - g_t(Y_t))\right).$$

• Define the level- $\alpha$  test:  $\tau = \min \{ t \ge 1 : W_t \ge \frac{1}{\alpha} \}$ 

#### Summary: Steps of our sequential test

#### Initialization:

- $\cdot$  A function class  ${\cal G}$
- a prediction strategy  $(A_P)$  to select  $\{g_t : t \ge 1\}$
- ONS betting strategy  $(\mathcal{A}_B)$  to select  $\{\lambda_t : t \geq 1\}$
- $W_0 = 1$

For t = 1, 2, ...:

- Get the next  $g_t$  from the prediction strategy,  $\mathcal{A}_{P}$ .
- Get the next  $\lambda_t$  from the betting strategy,  $\mathcal{A}_{\scriptscriptstyle B}$ .
- Observe the next pair  $(X_t, Y_t)$ .
- Update  $W_t = W_{t-1} \times (1 + \lambda_t (g_t(X_t) g_t(Y_t))).$
- Reject  $H_0$ , if  $W_t \ge 1/\alpha$ .

#### **Performance Guarantees**

Smaller Regret of  $\mathcal{A}_P \Rightarrow$  Faster growth of  $W_t \Rightarrow$  Stronger properties of the test  $\tau$ .

#### Performance Guarantees

#### Regret of $\mathcal{A}_P$

$$\mathcal{R}_n(\mathcal{A}_P) = \sup_{g \in \mathcal{G}} \left[ \left( \sum_{t=1}^n g(X_t) - g(Y_t) \right) - \left( \sum_{t=1}^n g_t(X_t) - g_t(Y_t) \right) \right].$$

#### Performance Guarantees

#### Regret of $\mathcal{A}_{P}$

$$\mathcal{R}_n(\mathcal{A}_P) = \sup_{g \in \mathcal{G}} \left[ \left( \sum_{t=1}^n g(X_t) - g(Y_t) \right) - \left( \sum_{t=1}^n g_t(X_t) - g_t(Y_t) \right) \right].$$

#### Regret-Power Connections under **H**<sub>1</sub> (Informal)

- If  $\lim_{n\to\infty} \mathcal{R}_n/n$  is smaller than  $d_{\mathcal{G}}(P_X, P_Y)$  w.p. 1, the test  $\tau$  is consistent.
- If  $\mathcal{R}_n/n \to 0$  with sufficiently large probability, then  $\mathbb{E}[\tau]$  is finite.
- If the  $\mathcal{R}_n/n \to 0$  w.p. 1, then the test  $\tau$  is exponentially consistent.

#### **Application 1: Sequential KS Test**

• 
$$\mathcal{X} = \mathbb{R}$$
, and  $\mathcal{G} = \{\mathbf{1}_{(-\infty,u]} - 0.5 : u \in \mathbb{R}\}.$ 

• Plug-in prediction strategy ( $A_{plug-in}$ ):  $g_t = \mathbf{1}_{(-\infty,u_t]} - 0.5$ , where

$$u_t \in \underset{u \in \mathbb{R}}{\operatorname{arg\,max}} \ \widehat{F}_{X,t-1}(u) - \widehat{F}_{Y,t-1}(u).$$

- For any  $n \ge 1$ ,  $\mathcal{R}_n(\mathcal{A}_{plug-in})/n = O(1/\sqrt{n})$ , w.p.  $1 1/n^2$ .
- Hence, the resulting test is consistent, and satisfies  $\mathbb{E}[\tau] = \mathcal{O}\left(1/d_{\text{KS}}^2(P_X, P_Y)\right)$  under  $H_1$ .
- There exist distributions on which any test,  $\tau'$ , must have  $\mathbb{E}[\tau'] = \Omega\left(1/d_{KS}^2(P_X, P_Y)\right).$

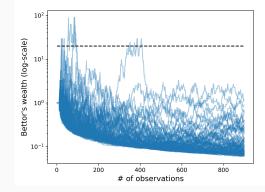
#### **Application 2: Sequential Kernel MMD Test**

- General  $\mathcal{X}$ , and  $\mathcal{G} = \{g \in \mathsf{RKHS}(k) : ||g||_k \le 1\}.$
- + Projected Gradient Ascent prediction strategy ( $\mathcal{A}_{\text{PGA}})$
- $\mathcal{A}_{PGA}$  satisfies  $\mathcal{R}_n(\mathcal{A}_{PGA})/n = \mathcal{O}(1/\sqrt{n})$ , w.p. 1.
- Hence, the resulting test is exponentially consistent, and satisfies  $\mathbb{E}[\tau] = \mathcal{O}\left(1/d_{\text{MMD}}^2(P_X, P_Y)\right)$  under  $H_1$ .
- There exist distributions on which any test,  $\tau'$ , satisfies  $\mathbb{E}[\tau'] = \Omega \left( 1/d_{\text{MMD}}^2(P_X, P_Y) \right)$ , under  $H_1$ .

#### An Example

**Under**  $H_0$ , the wealth process of the bettor rarely exceeds the level  $1/\alpha$ .

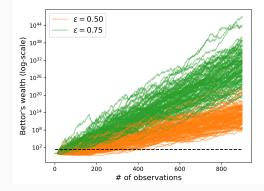
$$P_X = N(0, 1)$$
  
 $P_Y = N(0, 1)$ 



#### An Example

**Under**  $H_1$ , the wealth process with the plug-in prediction strategy, grows at an exponential rate.

$$P_X = N(0, 1)$$
$$P_Y = N(\varepsilon, 1)$$



# Extension to time-varying distributions

Our ideas easily extend to the following case:

For t = 1, 2, ...:

- Bettor selects  $g_t$  and  $\lambda_t$ .
- Adversary selects distributions  $P_{X,t}$  and  $P_{Y,t}$ .
- The pair,  $(X_t, Y_t) \sim P_{X,t} \times P_{Y,t}$  is revealed.
- Update the wealth:  $W_t = W_{t-1} \times (1 + \lambda_t (g_t(X_t) g_t(Y_t))).$
- Reject the null if  $W_t \ge 1/\alpha$ .

Under some mild assumptions on  $\mathcal{G}$ , the test defined above is consistent.

- Relaxing the assumption of paired observations.
- $\cdot\,$  Relaxing the boundedness assumption on the functions in  $\mathcal{G}.$
- A general problem unifying several tasks such as two-sample testing, independence testing, and symmetry testing.

# Thank you.

• 
$$\limsup_{n\to\infty} \frac{\mathcal{R}_n}{n} < d_{\mathcal{G}}(P_X, P_Y) \text{ a.s. } \Rightarrow \mathbb{P}_{P_{XY}}(\tau < \infty) = 1.$$

• For a sequence  $r_n \rightarrow 0$ , define  $E_n = \{\mathcal{R}_n / n \leq r_n\}$ . Then,

$$\sum_{n\geq 1}\mathbb{P}_{P_{XY}}(E_n^c)<\infty \ \Rightarrow \ \mathbb{E}_{P_{XY}}[\tau]<\infty.$$

• If  $\mathbb{P}_{P_{XY}}(E_n^c) = 0$  for some  $r_n \to 0$ , then we have

$$\liminf_{n\to\infty}\frac{-1}{2n}\mathbb{P}_{P_{XY}}(\tau>n)=\beta^*.$$
 (optimal exponent)

#### Testing invariance to an operator

- Given a stream of observations:  $U_1, U_2, \ldots$  on  $\mathcal{U}$ , drawn i.i.d.from  $P_{U}$ .
- Let  $T: \mathcal{U} \to \mathcal{U}$  be a known operator.
- Consider the problem:

 $H_0: P_U = P_U \circ T^{-1}$ , versus  $H_1: P_U \neq P_U \circ T^{-1}$ .

- This formulation unifies several problems such as two-sample testing, independence testing, and symmetry testing.
- For two-sample testing:

$$\mathcal{U} = \mathcal{X} \times \mathcal{X}, \quad \mathcal{U} = (X, Y), \quad P_U = P_X \times P_Y$$
  
 $T : \mathcal{X} \times \mathcal{X} \to \mathcal{X} \times \mathcal{X}, \quad \text{such that} \quad T(x, y) = (y, x).$